

Eluder Dimension and the Sample Complexity of Optimistic Exploration

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JULY 22, 2020

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Problem formulation

- A set of actions \mathcal{A} .
- A set of real-valued functions $\mathcal{F} = \{f_\rho : \mathcal{A} \mapsto \mathbb{R} \mid \rho \in \Theta\}$.
- At each time t , the agent is presented with a subset $\mathcal{A}_t \subset \mathcal{A}$.
- The agent selects an action $A_t \in \mathcal{A}_t$, and then receive a reward R_t .
- H_t is the history $(\mathcal{A}_1, A_1, R_1, \dots, \mathcal{A}_{t-1}, R_{t-1}, \mathcal{A}_t)$.
- The agent employs a policy $\pi = \{\pi_t \mid t \in \mathbb{N}\}$, $\pi_t(H_t)$ is a distribution over \mathcal{A} with support \mathcal{A}_t .
- We assume that $\mathbb{E}[R_t \mid H_T, \theta, A_t] = f_\theta(A_t)$.

Problem formulation

- The T-proiod regret of a policy π is defined by

$$\text{Regret}(T, \pi, \theta) = \sum_{t=1}^T \mathbb{E}[\max_{a \in \mathcal{A}_t} f_{\theta}(a) - f_{\theta}(A_t) | \theta].$$

- The T-period Bayesian regret is defined by $\mathbb{E}[\text{Regret}(T, \pi, \theta)]$, where the expectation is taken with respect to the prior distribution over θ . Hence,

$$\text{BayesRegret}(T, \pi) = \sum_{t=1}^T \mathbb{E}[\max_{a \in \mathcal{A}_t} f_{\theta}(a) - f_{\theta}(A_t)].$$

VC dimension

Given a sample $\mathcal{D}_n = \{(x_1, y_1), \dots, (x_n, y_n)\}$, and define $S = \{x_1, \dots, x_n\}$. Consider the set

$$\mathcal{H}_S = \mathcal{H}_{x_1, \dots, x_n} = \{(h(x_1), \dots, h(x_n)) : h \in \mathcal{H}\}.$$

Definition (Growth Function). *The growth function is the maximum number of ways into which n points can be classified by the function class:*

$$G_{\mathcal{H}}(n) = \sup_{x_1, \dots, x_n} |\mathcal{H}_S|.$$

Definition (VC Dimension). *The VC dimension of a class \mathcal{H} is the largest $n = d_{VC}(\mathcal{H})$ such that*

$$G_{\mathcal{H}}(n) = 2^n.$$

In other words, VC dimension of a function class \mathcal{H} is the cardinality of the largest set that it can shatter.

Example

- A finite binary-valued function class
 $\mathcal{F} = \{f_\rho : \mathcal{A} \mapsto \{0, 1\} \mid \rho \in \{1, \dots, n\}\}$.
- A finite action set $\mathcal{A} = \{1, \dots, n\}$.
- Let $f_\rho(a) = \mathbf{1}(\rho = a)$.
- In time step t , $R_t = f_\rho(A_t)$.
- If ρ is uniformly distributed over $\{1, \dots, n\}$, it's easy to see that the Bayesian regret grows linearly with n .

We formulate this problem as a supervised learning problem:

- At each time step, an action A_t is sampled uniformly from \mathcal{A} and the reward $f_\theta(A_t)$ is observed.
- For large n , the time it takes to effectively learn to predict $f_\theta(A_t)$.

Eluder dimension

DEFINITION 2. An action $a \in \mathcal{A}$ is ϵ -dependent on actions $\{a_1, \dots, a_n\} \subseteq \mathcal{A}$ with respect to \mathcal{F} if any pair of functions $f, \tilde{f} \in \mathcal{F}$, satisfying $\sqrt{\sum_{i=1}^n (f(a_i) - \tilde{f}(a_i))^2} \leq \epsilon$ also satisfies $f(a) - \tilde{f}(a) \leq \epsilon$. Further, a is ϵ -independent of $\{a_1, \dots, a_n\}$ with respect to \mathcal{F} if a is not ϵ -dependent on $\{a_1, \dots, a_n\}$.

DEFINITION 3. The ϵ -eluder dimension $\dim_E(\mathcal{F}, \epsilon)$ is the length d of the longest sequence of elements in \mathcal{A} such that, for some $\epsilon' \geq \epsilon$, every element is ϵ' -independent of its predecessors.

DEFINITION 4. An action a is *VC-independent* of $\tilde{\mathcal{A}} \subseteq \mathcal{A}$ if for any $f, \tilde{f} \in \mathcal{F}$, there exists some $\tilde{f} \in \mathcal{F}$, which agrees with f on a and with \tilde{f} on $\tilde{\mathcal{A}}$; that is, $\tilde{f}(a) = f(a)$ and $\tilde{f}(\tilde{a}) = \tilde{f}(\tilde{a})$ for all $\tilde{a} \in \tilde{\mathcal{A}}$. Otherwise, a is *VC-dependent* on $\tilde{\mathcal{A}}$.

DEFINITION 5. The VC dimension of a class of binary-valued functions with domain \mathcal{A} is the largest cardinality of a set $\tilde{\mathcal{A}} \subseteq \mathcal{A}$ such that every $a \in \mathcal{A}$ is VC-independent of $\tilde{\mathcal{A}} \setminus \{a\}$.

Eluder dimension for common function class

- Finite action spaces.

For all $\epsilon > 0$, the ϵ -eluder dimension of \mathcal{A} is bounded by $|\mathcal{A}|$.

- Euclid space.

The dimension of Euclid space for linear functions is 0-eluder dimension of this space.

- Reward functions are parameterized by a vector $\theta \in \Theta \subset \mathbb{R}^d$.
- Feature mapping $\phi : \mathcal{A} \mapsto \mathbb{R}^d$ such that $f_\theta(a) = \langle \phi(a), \theta \rangle$.
- An ellipsoidal confidence set $\Theta_t = \{\rho \in \mathbb{R}^d : \|\rho - \hat{\theta}_t\|_{V_t} \leq \sqrt{\beta_t}\}$, where $V_t := \sum_{k=1}^t \phi(A_k)\phi(A_k)^T + \lambda I$ for some $\lambda \in \mathbb{R}$ captures the amount of exploration carried out in each direction up to time t .

Algorithm 2 (Linear-Gaussian UCB)

1. Update Statistics:

$$\mu_t \leftarrow \mathbb{E}[\theta \mid H_t]$$

$$\Sigma_t \leftarrow \mathbb{E}[(\theta - \mu_t)(\theta - \mu_t)^\top \mid H_t]$$

2. Select Action:

$$\bar{A}_t \in \arg \max_{a \in \mathcal{A}} \{ \langle \phi(a), \mu_t \rangle + \beta \log(t) \|\phi(a)\|_{\Sigma_t} \}$$

3. Increment t and Go to Step 1.

Theorem

Assume there exist constants γ and S such that for all $a \in \mathcal{A}$ and $\rho \in \Theta$, $\|\rho\|_2 \leq S$ and $\|\phi(a)\|_2 \leq \gamma$. Then

$$\dim_E(\mathcal{F}, \epsilon) \leq 3d(e/(e-1)) \log\{3 + 3((2S)/\epsilon)^2\} + 1.$$

Define

$$\omega_k := \sup \left\{ (f_{\rho_1} - f_{\rho_2})(a_k) : \sqrt{\sum_{i=1}^{k-1} (f_{\rho_1} - f_{\rho_2})^2(a_i)} \leq \epsilon' \rho_1, \rho_2 \in \Theta \right\}.$$

Then the ϵ -eluder dimension is the longest sequence such that $\omega_k > \epsilon'$.

- Let $\phi_k = \phi(a_k)$,
- $\rho = \rho_1 - \rho_2$,
- and $\Phi_k = \sum_{i=1}^{k-1} \phi_i \phi_i^T$.
- In this case, $\sum_{i=1}^{k-1} (f_{\rho_1} - f_{\rho_2})^2(a_i) = \rho^T \Phi_k \rho$,
- and by the triangle inequality $\|\rho\|_2 \leq 2S$.

Step 1. If $\omega_k \geq \epsilon'$, then $\phi_k^T V_k^{-1} \phi_k \geq \frac{1}{2}$, where $V_k := \Phi_k + \lambda I$ and $\lambda = (\epsilon' / (2S))^2$.

Proof.

$$\begin{aligned}\omega_k &\leq \max\{\rho^T \phi_k : \rho^T \Phi_k \rho \leq (\epsilon')^2, \rho^T I \rho \leq (2S)^2\} \\ &\leq \max\{\rho^T \phi_k : \rho^T V_k \rho \leq 2(\epsilon')^2\} \\ &= \sqrt{2(\epsilon')^2} \|\phi_k\|_{V_k^{-1}}\end{aligned}$$



Step 2. If $\omega_i \geq \epsilon'$, for each $i < k$, then $\det V_k \geq \lambda^d \left(\frac{3}{2}\right)^{k-1}$ and $\det V_k \leq ((\gamma^2(k-1))/d + \lambda)^d$.

Proof.

PROOF. Since $V_k = V_{k-1} + \phi_k \phi_k^T$, using the matrix determinant lemma,

$$\det V_k = \det V_{k-1} (1 + \phi_k^T V_{k-1}^{-1} \phi_k) \geq \det V_{k-1} \left(\frac{3}{2}\right) \geq \dots \geq \det[\lambda I] \left(\frac{3}{2}\right)^{k-1} = \lambda^d \left(\frac{3}{2}\right)^{k-1}.$$

Recall that $\det V_k$ is the product of the eigenvalues of V_k , whereas $\text{trace}[V_k]$ is the sum. As noted in Dani et al. [16], $\det V_k$ is maximized when all eigenvalues are equal. This implies $\det V_k \leq ((\text{trace}[V_k])/d)^d \leq ((\gamma^2(k-1))/d + \lambda)^d$. \square



Step 3. Complete proof.

Proof.

Step 2 shows that $(\frac{3}{2})^{(k-1)d} \leq \alpha_0[(k-1)/d] + 1$, where $\alpha_0 = \gamma^2/\lambda = (2S\gamma/\epsilon')^2$.

Let $B(x, \alpha) = \max\{B : (1+x)^B \leq \alpha B + 1\}$, then the number of times $\omega_k > \epsilon'$ can occur is bounded by $dB(1/2, \alpha_0) + 1$.

We now derive an explicit bound on $B(x, \alpha)$ for any $x \leq 1$. Note that any $B \geq 1$ must satisfy the inequality $\ln\{1+x\}B \leq \ln\{1+\alpha\} + \ln B$. Since $\ln\{1+x\} \geq x/(1+x)$, using the transformation of variables $y = B[x/(1+x)]$ gives

$$y \leq \ln\{1+\alpha\} + \ln \frac{1+x}{x} + \ln y \leq \ln\{1+\alpha\} + \ln \frac{1+x}{x} + \frac{y}{e} \implies y \leq \frac{e}{e-1} \left(\ln\{1+\alpha\} + \ln \frac{1+x}{x} \right).$$

This implies $B(x, \alpha) \leq ((1+x)/x)(e/(e-1))(\ln\{1+\alpha\} + \ln((1+x)/x))$. The claim follows by plugging in $\alpha = \alpha_0$ and $x = 1/2$. \square



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Algorithm 3 (Independent posterior sampling)

1. **Sample Model:**

$$\hat{\theta}_t \sim N(\mu_{t-1}, \Sigma_{t-1})$$

2. **Select Action:**

$$A_t \in \arg \max_{a \in \mathcal{A}} \hat{\theta}_t(a)$$

3. **Update Statistics:** For each a ,

$$\mu_{ta} \leftarrow \mathbb{E}[\theta_a | H_t]$$

$$\Sigma_{taa} \leftarrow \mathbb{E}[(\theta_a - \mu_{ta})^2 | H_t]$$

4. **Increment t and Go to Step 1.**

Algorithm 4 (Linear posterior sampling)

1. **Sample Model:**

$$\hat{\theta}_t \sim N(\mu_{t-1}, \Sigma_{t-1})$$

2. **Select Action:**

$$A_t \in \arg \max_{a \in \mathcal{A}} \langle \phi(a), \hat{\theta}_t \rangle$$

3. **Update Statistics:**

$$\mu_t \leftarrow \mathbb{E}[\theta | H_t]$$

$$\Sigma_t \leftarrow \mathbb{E}[(\theta - \mu_t)(\theta - \mu_t)^\top | H_t]$$

4. **Increment t and Go to Step 1.**

UCB regret decomposition

Consider a UCB algorithm with a UCB sequence $U = \{U_t | t \in \mathbb{N}\}$. Let $\bar{A}_t \in \arg \max_{a \in \mathcal{A}_t} U_t(a)$ and $A^* \in \arg \max_{a \in \mathcal{A}_t} f_\theta(a)$. We have the following regret decomposition:

$$\begin{aligned} f_\theta(A_t^*) - f_\theta(\bar{A}_t) &= f_\theta(A_t^*) - U_t(\bar{A}_t) + U_t(\bar{A}_t) - f_\theta(\bar{A}_t) \\ &\leq [f_\theta(A_t^*) - U_t(A_t^*)] + [U_t(\bar{A}_t) - f_\theta(\bar{A}_t)] \end{aligned}$$

Bayesian regret:

$$\text{BayesRegret}(T, \pi^U) \leq \mathbb{E} \sum_{t=1}^T [U_t(\bar{A}_t) - f_\theta(\bar{A}_t)] + \mathbb{E} \sum_{t=1}^T [f_\theta(A_t^*) - U_t(A_t^*)]$$

Posterior sampling regret decomposition

PROPOSITION 1. For any UCB sequence $\{U_t \mid t \in \mathbb{N}\}$,

$$\text{BayesRegret}(T, \pi^{PS}) = \mathbb{E} \sum_{t=1}^T [U_t(A_t) - f_{\theta}(A_t)] + \mathbb{E} \sum_{t=1}^T [f_{\theta}(A_t^*) - U_t(A_t^*)]$$

for all $T \in \mathbb{N}$.

Posterior sampling regret decomposition

PROOF. Note that, conditioned on H_t , the optimal action A_t^* and the action A_t selected by posterior sampling are identically distributed, and U_t is deterministic. Hence $\mathbb{E}[U_t(A_t^*) | H_t] = \mathbb{E}[U_t(A_t) | H_t]$. Therefore

$$\begin{aligned}\mathbb{E}[f_\theta(A_t^*) - f_\theta(A_t)] &= \mathbb{E}[\mathbb{E}[f_\theta(A_t^*) - f_\theta(A_t) | H_t]] \\ &= \mathbb{E}[\mathbb{E}[U_t(A_t) - U_t(A_t^*) + f_\theta(A_t^*) - f_\theta(A_t) | H_t]] \\ &= \mathbb{E}[\mathbb{E}[U_t(A_t) - f_\theta(A_t) | H_t] + \mathbb{E}[f_\theta(A_t^*) - U_t(A_t^*) | H_t]] \\ &= \mathbb{E}[U_t(A_t) - f_\theta(A_t)] + \mathbb{E}[f_\theta(A_t^*) - U_t(A_t^*)].\end{aligned}$$

Summing over t gives the result. \square

- Assumption 1. For all $f \in \mathcal{F}$ and $a \in \mathcal{A}$, $f(a) \in [0, C]$.
- Assumption 2. For all $t \in \mathbb{N}$, $R_t - f\theta(A_t)$ conditioned on (H_t, θ, A_t) is σ -sub-Gaussian.

Let $L_{2,t}(f) = \sum_1^{t-1} (f(A_t) - R_t)^2$, and $\hat{f}_t^{LS} \in \arg \min_{f \in \mathcal{F}} L_{2,t}(f)$.

LEMMA 3. For any $\delta > 0$ and $f: \mathcal{A} \mapsto \mathbb{R}$, with probability at least $1 - \delta$,

$$L_{2,t}(f) \geq L_{2,t}(f_\theta) + \frac{1}{2} \|f - f_\theta\|_{2,E_t}^2 - 4\sigma^2 \log(1/\delta)$$

simultaneously for all $t \in \mathbb{N}$.

$$\beta_t^*(\mathcal{F}, \delta, \alpha) := 8\sigma^2 \log(N(\mathcal{F}, \alpha, \|\cdot\|_\infty)/\delta) + 2\alpha t(8C + \sqrt{8\sigma^2 \ln(4t^2/\delta)}).$$

PROPOSITION 6. For all $\delta > 0$ and $\alpha > 0$, if

$$\mathcal{F}_t = \{f \in \mathcal{F} : \|f - \hat{f}_t^{LS}\|_{2, E_t} \leq \sqrt{\beta_t^*(\mathcal{F}, \delta, \alpha)}\}$$

for all $t \in \mathbb{N}$, then

$$\mathbb{P}\left(f_\theta \in \bigcap_{t=1}^{\infty} \mathcal{F}_t\right) \geq 1 - 2\delta.$$

DEFINITION 1. The *Kolmogorov dimension* of a function class \mathcal{F} is given by

$$\dim_K(\mathcal{F}) = \limsup_{\alpha \downarrow 0} \frac{\log N(\mathcal{F}, \alpha, \|\cdot\|_\infty)}{\log(1/\alpha)}.$$

In particular, we have the following result.

PROPOSITION 7. For any fixed class of functions \mathcal{F} ,

$$\beta_t^*(\mathcal{F}, 1/t^2, 1/t^2) = 16(1 + o(1) + \dim_K(\mathcal{F})) \log t.$$

Bayesian regret bounds

LEMMA 4. For all $T \in \mathbb{N}$, if $\inf_{\rho \in \mathcal{F}_\tau} f_\rho(a) \leq f_\theta(a) \leq \sup_{\rho \in \mathcal{F}_\tau} f_\rho(a)$ for all $\tau \in \mathbb{N}$ and $a \in \mathcal{A}$ with probability at least $1 - 1/T$, then

$$\text{BayesRegret}(T, \pi^{PS}) \leq C + \mathbb{E} \sum_{t=1}^T w_{\mathcal{F}_t}(A_t).$$

Bayesian regret bounds

PROPOSITION 8. If $(\beta_t \geq 0 \mid t \in \mathbb{N})$ is a nondecreasing sequence and $\mathcal{F}_t := \{f \in \mathcal{F} : \|f - \hat{f}_t^{LS}\|_{2, E_t} \leq \sqrt{\beta_t}\}$, then

$$\sum_{t=1}^T \mathbf{1}(w_{\mathcal{F}_t}(A_t) > \epsilon) \leq \left(\frac{4\beta_T}{\epsilon^2} + 1 \right) \dim_E(\mathcal{F}, \epsilon)$$

for all $T \in \mathbb{N}$ and $\epsilon > 0$.

Bayesian regret bounds

LEMMA 5. If $(\beta_t \geq 0 \mid t \in \mathbb{N})$ is a nondecreasing sequence and $\mathcal{F}_t := \{f \in \mathcal{F} : \|f - \hat{f}_t^{LS}\|_{2, E_t} \leq \sqrt{\beta_t}\}$, then

$$\sum_{t=1}^T w_{\mathcal{F}_t}(A_t) \leq 1 + \dim_E(\mathcal{F}, T^{-1})C + 4\sqrt{\dim_E(\mathcal{F}, T^{-1})\beta_T T}$$

for all $T \in \mathbb{N}$.

PROPOSITION 10. For any fixed class of functions \mathcal{F} ,

$$\text{BayesRegret}(T, \pi^{PS}) \leq 1 + [\dim_E(\mathcal{F}, T^{-1}) + 1]C + 16\sigma\sqrt{\dim_E(\mathcal{F}, T^{-1})(1 + o(1) + \dim_K(\mathcal{F})) \log(T)T}$$

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PROPOSITION 8. If $(\beta_t \geq 0 \mid t \in \mathbb{N})$ is a nondecreasing sequence and $\mathcal{F}_t := \{f \in \mathcal{F} : \|f - \hat{f}_t^{LS}\|_{2, E_t} \leq \sqrt{\beta_t}\}$, then

$$\sum_{t=1}^T \mathbf{1}(w_{\mathcal{F}_t}(A_t) > \epsilon) \leq \left(\frac{4\beta_T}{\epsilon^2} + 1 \right) \dim_E(\mathcal{F}, \epsilon)$$

for all $T \in \mathbb{N}$ and $\epsilon > 0$.

Step 1: If $\omega_t(A_t) > \epsilon$, then A_t is ϵ -dependent on fewer than $4\beta_T/\epsilon^2$ disjoint subsequences of (A_1, \dots, A_{t-1}) for $T > t$.

$$\omega_t(A_t) > \epsilon \Rightarrow \exists \bar{f}, \underline{f} \in \mathcal{F}_t, \bar{f}(A_t) - \underline{f}(A_t) > \epsilon$$

\Rightarrow If A_t is ϵ -dependent on a subsequence of (A_1, \dots, A_{t-1}) ,

$$\text{then } \sum_{j=1}^k (\bar{f}(A_j) - \underline{f}(A_j))^2 > \epsilon^2$$

\Rightarrow If A_t is ϵ -dependent on K disjoint subsequences

$$\text{then } \|\bar{f} - \underline{f}\|_{2, E_t}^2 > K\epsilon^2$$

$$\|\bar{f} - \underline{f}\|_{2, E_t} \leq \|\bar{f} - \hat{f}_t^{LS}\|_{2, E_t} + \|\underline{f} - \hat{f}_t^{LS}\|_{2, E_t} \leq 2\sqrt{\beta_t} \leq 2\sqrt{\beta_T}$$

It's follows that $K < 4\beta_T/\epsilon^2$.

Step 2: For any action sequence (a_1, \dots, a_τ) , there is some element a_j that is ϵ -dependent on at least $\tau/d - 1$ disjoint subsequences, where $d := \dim_E(\mathcal{F}, \epsilon)$.

Step 3:

Consider taking (a_1, \dots, a_τ) to be the subsequence of (A_1, \dots, A_T) , in which the elements satisfies $\omega_t(A_t) > \epsilon$.

$$\Rightarrow \tau/d - 1 \leq 4\beta_T/\epsilon^2.$$

LEMMA 5. If $(\beta_t \geq 0 \mid t \in \mathbb{N})$ is a nondecreasing sequence and $\mathcal{F}_t := \{f \in \mathcal{F} : \|f - \hat{f}_t^{LS}\|_{2, E_t} \leq \sqrt{\beta_t}\}$, then

$$\sum_{t=1}^T w_{\mathcal{F}_t}(A_t) \leq 1 + \dim_E(\mathcal{F}, T^{-1})C + 4\sqrt{\dim_E(\mathcal{F}, T^{-1})\beta_T T}$$

for all $T \in \mathbb{N}$.

Step 1:

PROOF. To reduce notation, write $d = \dim_E(\mathcal{F}, T^{-1})$ and $w_t = w_t(A_t)$. Reorder the sequence $(w_1, \dots, w_T) \rightarrow (w_{i_1}, \dots, w_{i_T})$, where $w_{i_1} \geq w_{i_2} \geq \dots \geq w_{i_T}$. We have

$$\sum_{t=1}^T w_{\mathcal{F}_t}(A_t) = \sum_{t=1}^T w_{i_t} = \sum_{t=1}^T w_{i_t} \mathbf{1}\{w_{i_t} \leq T^{-1}\} + \sum_{t=1}^T w_{i_t} \mathbf{1}\{w_{i_t} > T^{-1}\} \leq 1 + \sum_{t=1}^T w_{i_t} \mathbf{1}\{w_{i_t} \geq T^{-1}\}.$$

Step 2:

- $d = \dim_E(\mathcal{F}, T^{-1}) \geq \dim_E(\mathcal{F}, \epsilon)$, for all $\epsilon > T^{-1}$.
- $\sum_{t=1}^T \mathbf{1}(\omega_t > \epsilon) < ((4\beta_T)/\epsilon^2 + 1)d$.
- $\omega_t > \epsilon \Rightarrow \sum_{k=1}^T \mathbf{1}(\omega_k > \epsilon) \geq t$.
- $\epsilon < \sqrt{(4\beta_T d)/(t - d)}$.
- $w_t \leq \min\{C, \sqrt{(4\beta_T d)/(t - d)}\}$.

Step 3:

$$\sum_{i=1}^T w_i \mathbf{1}\{w_i > T^{-1}\} \leq dC + \sum_{t=d+1}^T \sqrt{\frac{4d\beta_T}{t-d}} \leq dC + 2\sqrt{d\beta_T} \int_{t=0}^T \frac{1}{\sqrt{t}} dt = dC + 4\sqrt{d\beta_T T}.$$

Thanks!